

MEASUREMENTS AND CONCEPTS OF THE INEQUALITY
OF INCOMES:
AN ALTERNATIVE BEING STUDIED IN BRAZIL

By

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Thesis under the direction of Professor Samuel Morley

The main purpose of this work should be understood as a preliminary attempt to show that, in general, measures of inequality of income imply their own concepts of the process of income distribution between individuals or groups. These underlying concepts should be brought out through a non-conventional analysis of the currently used measures and their mathematical-statistical properties. In our view, this should become the centre of attention of researchers who are interested in understanding the causes and effects of inequality of income in a given economic system.

One should emphasize that the conceptual link between the measures of inequality and their subsequent socio-economic interpretations must exist in order to obtain better theoretical and empirical use of research on income distribution. Consequently a knowledge of the characteristics and properties of the measures of inequality of income makes it possible not only to choose those which are most appropriate to the theoretical models, but also to supply valuable additional elements for a better utilization in empirical analysis. This implies that these measures, so frequently used, are not limited only to furnishing quantitative estimates of inequality of income; in addition to this, being developed as they are on their own concepts of the process of income distribution, they allow the empirical verification and proof of these concepts by means of

appropriate manipulation of their operational characteristics and properties.

In this context the characteristics and properties of three well known measures of inequality of income are presented and commented on in the first part of this work; these are Pareto's index, Gini's coefficient of concentration and the variance of logarithms. Their respective concepts of the process of income distribution are brought out and evaluated in accordance with our interpretations and observations of other works. In the second part, taking as a basis a summary analysis of classical economic theory, an alternative measure on inequality of income is proposed. Its concept, mathematical formulation, operative statistical form as well as its fundamental characteristics and properties are developed. In the third part, the proposed measure is applied to estimate the inequality of income in Brazil in two specific cases: (a) Individual income distribution 1960/1970 and (b) Income distribution in the formal urban sector 1970/1973. Finally, the fourth part is directed to presenting the most important conclusions and suggestions for future works in this field.

Approved _____

Date _____

Adviser

... a man... can stare stupidly at phenomena; but in the
absence of imagination they will not connect themselves in
any rational way. (Italics mine.)

C. S. Pierce

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Obviously, I accept full blame for any remaining errors or omissions.

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TABLE OF CONTENTS

| | Page |
|--|------|
| ACKNOWLEDGMENTS..... | ii |
| LIST OF TABLES..... | v |
| LIST OF ILLUSTRATIONS..... | vi |
| INTRODUCTION..... | 1 |
| Chapter | |
| I. CONCEPTS OF THE DISTRIBUTION OF INCOME..... | 5 |
| Pareto's Index..... | 5 |
| Gini's Coefficient of Concentration..... | 9 |
| Variance of Logarithms..... | 12 |
| Further Observations..... | 14 |
| II. AN ALTERNATIVE CONCEPT..... | 16 |
| The Means of Production and Income Distribution..... | 16 |
| The Question of Interdependence..... | 18 |
| The Interpretation of the Process of Income Distribution..... | 19 |
| An Operational Form of the Interpretation of the Process of Income Distribution.... | 21 |
| Practical Application of the Operation... | 22 |
| The Range of Variation of the Measure.... | 25 |
| Graphical Presentation..... | 26 |
| Disaggregation of the Measure..... | 26 |
| Properties of the Measure..... | 33 |
| III. APPLICATION OF THE MEASURE..... | 36 |

| | Page |
|---|------|
| IV. SUMMARY, CONCLUSIONS AND SUGGESTIONS..... | 49 |
| APPENDIX..... | 52 |
| BIBLIOGRAPHY..... | 54 |

LIST OF TABLES

| Table | Page |
|--|------|
| 1. Distribution of Individual Incomes: Brazil 1960..... | 38 |
| 2. Distribution of Individual Incomes: Brazil 1970..... | 39 |
| 3. Disaggregation of $CV^2 / CV^2 + 1$ | 41 |
| 4. Distribution of Wages in the Formal Urban Sector by Regions, 1970..... | 43 |
| 5. Distribution of Wages in the Formal Urban Sector by Regions, 1971..... | 44 |
| 6. Distribution of Wages in the Formal Urban Sector by Regions, 1972..... | 45 |
| 7. Distribution of Wages in the Formal Urban Sector by Regions, 1973..... | 46 |

LIST OF ILLUSTRATIONS

| Figure | Page |
|--|------|
| 1. Pareto's Function..... | 7 |
| 2. Lorenz Curve..... | 10 |
| 3. Graphical Presentation: $CV^2/CV^2 + 1$ | 27 |
| 4. Disaggregation of $CV^2/CV^2 + 1$ | 42 |

INTRODUCTION

Considerable research has been carried out recently concerning income distribution both in theoretical approaches and empirical exercises. This interest is not merely academic but also political, to the extent that the results of these studies, inasmuch as they contribute to a better understanding of the phenomenon and its implications for the economic relationships in a given system of production, can lead to policy options for government programmes.

In the theoretical approaches, one notes as a general rule the "resurrection" of the classicists¹ and the improvement of the neo-classicists² in an attempt to grasp more clearly the causes and effects of the functioning of the economic system in the determination of income distribution.

In the empirical exercises we discover:

- (1) Analysis of the distribution of income of countries, regions, etc., through the application

¹See, for example, Maurice Dobb, "The Straffa System and Critique of the Neo-Classical Theory of Distribution", A Critique of Economic Theory, ed. by E.K. Hunt and Fesse G. Schwartz (Harmondsworth, Middlesex: Penguin Books, Inc., 1972), and Teoria del Valor y de la Distribución desde Adam Smith (Argentina: Siglo XXI Editores S.A., 1975); Joan Robinson, "Capital Theory Up to Date" and "Prelude to a Critique of Economic Theory", A Critique of Economic Theory, op. cit.

²Chiefly, Gary Becker, Human Capital (N.Y.: National Bureau of Economic Research, 1964) and Theodore W. Schultz, O Capital Humano (Rio de Janeiro: Zahar Editores, 1973).

of the models and measures available³.

(2) Works directed only at the perfection of the measurement of the inequality of income distribution; these usually start with the manipulation of mathematical/statistical relationships which provide operational instruments for estimating the inequality of incomes, with special references to particular features of the phenomenon of income distribution⁴.

(3) A further group of researchers try to analyse the "performance" of the measures of inequality of incomes. They are placed alongside each other to check the extent to which they estimate the "depth" and precision of income inequalities.

One begins to discover, on the basis of these works⁵, that suitable measurements can be developed for the kinds of hypotheses which should be tested in studies of income distri

³See, for example, Albert Fishlow, "Brazilian Size Distribution of Income", American Economic Review (May, 1972), 391-402; Carlos Geraldo Langoni, Distribuição da Renda e Desenvolvimento Econômico do Brasil (Rio de Janeiro: Editora Expressão e Cultura S.A., 1973); Rodolfo Hoffman and João Carlos Duarte, "A Distribuição de Renda no Brasil", Revista de Administração de Empresas (São Paulo), No. 2 (June, 1972), 46-66.

⁴See, for example, Anthony B. Atkinson, "On the Measurement of Inequality", Journal of Economic Theory, II (1970), 244-263; O. Elteto and E. Frigyes, "New Income Inequality Measures as Efficient Tools for Causal Analysis and Planning", Econometrica, XXXVI (April, 1968), 383-396.

⁵See, for example, Ramonaval Augusto Costa, "Medidas de Desigualdade de Renda", Boletim Geográfico (Rio de Janeiro), No. 238 (Jan./Feb., 1974), 45-72; João Carlos Duarte, "Aspectos da Distribuição da Renda no Brasil em 1970" (unpublished M.A. Thesis, University of Piracicaba, São Paulo, 1972); Rodolfo Hoffman, "Contribuição à Análise da Distribuição da Renda e da Posse de Terra no Brasil" (unpublished Livre Docência Dissertation, University of Piracicaba, São Paulo, 1971).

bution. Indirectly, one can also perceive that each measure of inequality of income brings with it, explicitly or implicitly, a concept of the operation of the process of income distribution even though this group of researchers does not explicitly mention this conceptual relationship. The link between the concept and the operation of the process becomes credible as a way of making more valid these measures and their respective analytical properties.

In view of these considerations, this thesis has three objectives. Firstly, to list some of the measures traditionally employed in estimating the inequality of income and attempt to evaluate their concepts of income distribution and their measurement properties. Secondly, to supply an alternative interpretation of the process through a further concept of the mechanism underlying the determination of the distribution of income generated by the productive process between groups of individuals. Thirdly, to operationalize an interpretation of the process through an alternative measure of inequality of income, applying it to a few income distributions available for Brazil.

Combining the first and second objectives, the intention is to provide systematic view of the analysis of income inequality based on a specific conceptual reference. With this, one may hope to warn the user that the consistent use of a measure of income inequality implies a certain conceptualization of the process of income distribution. It is hoped to have provided some elements which may remind researchers that, ultimately, underlying a certain measure of income distribution there is an understanding of how inequality of income comes about. This has tended to be forgotten in many of the studies in this field. The importance of the last objective, the operational aspect, is secondary to the extent to which one intends to understand the process as such, while making it clear that the improve

ment of studies on income distribution, especially the estimates of inequality, requires a systematic effort at understanding the features which involve the economic system.

CHAPTER I

CONCEPTS OF THE DISTRIBUTION OF INCOME

It is intended in this section to present in summary form, on the one hand how the phenomenon of the distribution of income is conceived by some of the measures of inequality of income and, on the other to develop some comments about the features of these measures.

Pareto's Index

Vilfredo Pareto⁶ introduced a measure of inequality of income which tried to show, in a simplified way, how the distribution of incomes manifests itself among individuals. His concept of the phenomenon can be said to be "empirical" insofar as it originated in the observations of the profiles of income distribution in various countries of his time.

The mathematical representation of his function of income distribution would be⁷

$$N = AY^{-b}$$

⁶Regarding Pareto's index see: Ramonaval Augusto Costa, "Medidas de Desigualdade de Renda", op.cit.; Rodolfo Hoffman, "Contribuição à Análise da Distribuição da Renda e da Posse de Terra no Brasil", op.cit.; Jorge Kingston, "A Desigualdade na Distribuição das Rendas", Revista Brasileira de Economia (Rio de Janeiro), No. 1 (March, 1952), 9-13.

⁷This is known as the first of Pareto's laws, that is the simplified form of his function of income distribution. Concerning Pareto's second law see: Rodolfo Hoffman, "Contribuição à Análise da Distribuição da Renda e da Posse de Terra no Brasil", op.cit.; Ramonaval Augusto Costa, "Medidas de Desigualdade de Renda", op.cit.

where N represents the number of people with income Y or more; A and b are constants. The constant b is called Pareto's coefficient (or index).

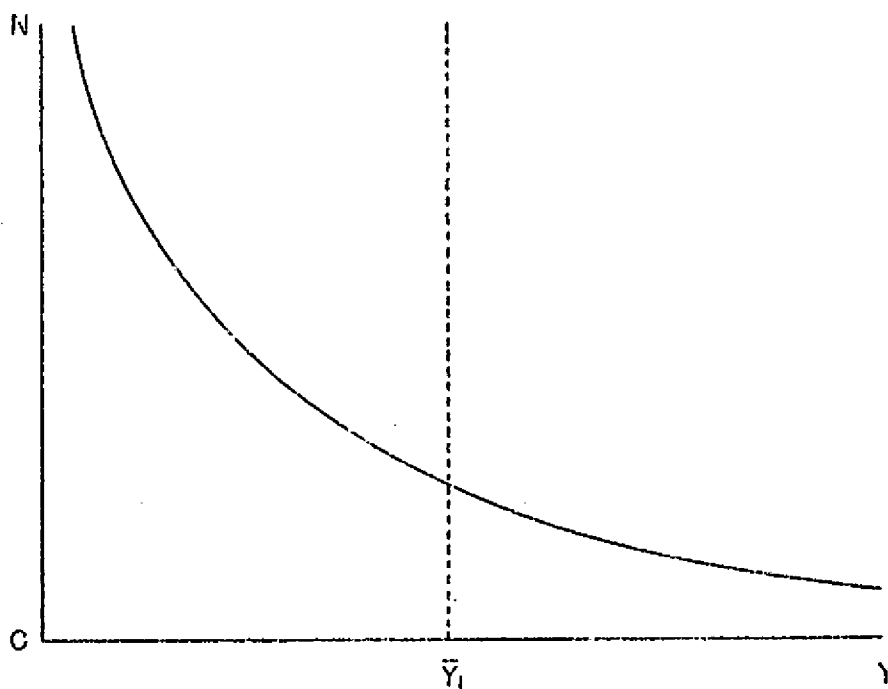
It will be noted from the interpretation of the Pareto function (in hyperbolic form) that the number of people with an income Y_i or more decreases as income increases, in other words the number of people with an income \bar{Y}_i or more is inversely proportional to the levels of income \bar{Y}_i . This inverse relationship is shown logically by the coefficient b , the differential marginal coefficient of the Pareto function. The graphical presentation of this statement can be seen in figure 1.

Some points may be underlined among the observations made about Pareto's index:

(1) The implicit concept of inequality of income among individuals (or groups of individuals) can be understood as the visual presentation of a social pyramid where the individuals at the base receive average incomes less than the individuals at the top. This inverse relationship between the number of individuals and the average levels of income would be, according to Pareto, the most common aspect of the distributions of income among the units observed (countries, regions, etc.) independently of the socio-economic conditions which generated these distributions. Consequently, his index represents the degree of inequality of income existing in a particular distribution, given the productive relationships preceding the distribution of the proceeds among the individuals.

(2) It can be said that his index reflects a "partial" concept of the phenomenon of distribution. "Partial" to the extent that it reflects

FIGURE 1
PARETO'S FUNCTION
($N = AY^{-b}$)



a simple functional relationship between individuals and the respective average levels of income without any attempt to perceive the mechanism of distribution of income as such.

(3) Some other obvious characteristics⁸ of Pareto's index separated into two groups are summarized briefly below:

(a) Advantages:

- i) Its calculation is relatively simple.
- ii) It is generally used for estimating the average incomes of the highest and lowest groups (when these are represented with the respective lower and upper limits left open).
- iii) Its use is immediate irrespective of the presentation of the distribution.

(b) Disadvantages:

- i) It is not very sensitive to small differences between the distributions of income.
- ii) Because of the fact that the adjustment of the curve to the observed data is only significant above a certain level of income \bar{Y}_1 (see figure 1), it often happens that only the upper part of the distribution is well interpreted by the Pareto function; this does not happen with the lower part of the distribution where the inequality of income (as a

⁸For better information on these characteristics, see: Ramonaval Augusto Costa, ibid.

general rule) is less, so that Pareto's reasoning is less applicable.

iii) A poor adjustment of the function to the observed data produces an inadequate interpretation of the concentration of income distribution through his index b.

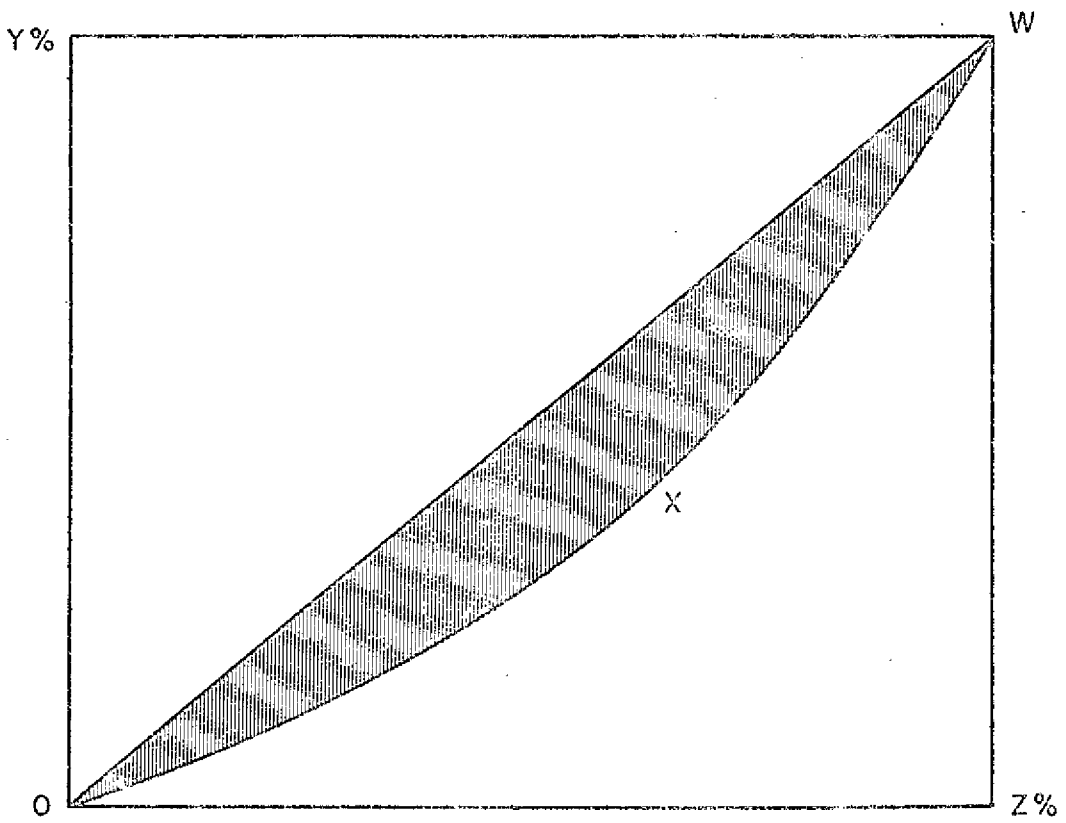
Gini's Coefficient of Concentration

Corrado Gini made a dual contribution to the study of the concentration of incomes. He developed two measures, one of them little known to researchers and described as Gini's index, and the other generally used, the Gini's coefficient of concentration. We are dealing here with the latter, not only because of its greater use by researchers, but also because of its practical significance when associated with the Lorenz curve.

Given the Lorenz curve, Gini's coefficient of concentration (G)⁹ can be visualised (see figure 2). One observes that the area OW (neutral area) refers to the situation of perfect equality of income (where all the individuals have the same share of total income); that the area OZW refers to the situation of extreme inequality of income (where only one individual receives the whole income of the distribution); and finally that the area OXW refers to the situation of intermediate inequality of income (where the shares of some individuals or groups in the total income exceed the shares of others in a differential manner).

⁹The Lorenz curve is obtained by placing on the horizontal axis the cumulative percentages of population and on the vertical axis the corresponding cumulative percentages of total income. In this way, one can visualise the participation in total income of each respective group of individuals being considered.

FIGURE 2
LORENZ CURVE



To obtain G through the Lorenz curve, it is sufficient to relate the areas OXW and OZW as follows:

$$G = \frac{OXW}{OZW}$$

whose limits of variance are 0 (area OW) and 1 (area OZW).

Given these initial observations, we may consider below the main aspects related to Gini's coefficient of concentration (G):

(1) The implicit concept of the inequality of income in G can be interpreted simply by the different income shares of individuals generated by the productive process. This relative differentiation therefore will reflect the degree of the inequality of income existing in the distribution.

(2) Compared with a simple functional relationship as in the case of Pareto's index, G shows in a clearer way the phenomenon of income distribution. Specifically a better or worse apparent concentration of income in a given distribution will depend directly on the share of total income of the individuals which produced it; the more unequal the shares, the worse is the degree of concentration of income in G and vice-versa.

(3) Summarized below are some characteristics of Gini's coefficient of concentration:¹⁰

(a) Advantages:

- i) It is sensitive to small variations between distributions of income.
- ii) It is a simple and clear concept (particularly when linked with the Lorenz curve).
- iii) It is easily and quickly calculated.

¹⁰See Ramonaval Augusto Costa, "Medidas de Desigualdade de Renda", op.cit.

- iv) It reflects the inequality of income in the whole of the distribution.
 - v) It can be disaggregated into components of concentration "within" and "between" distributions.
- (b) Disadvantages:
- i) Its underestimation of the area of concentration of income¹¹ by the methods of calculation traditionally used.
 - ii) The interpretation of its value becomes ambiguous in the case where two Lorenz curves intersect with the same area of concentration.

Variance of Logarithms

This deals basically with a measure of dispersion relative to the average income of the distribution, with the specification that one takes instead of absolute values the logarithms of the incomes (individuals or groups of individuals). The use of logarithms arises from the effort to minimize the effect of the absolute differences of income.

The expression for the variance of logarithms is the following:

$$\sigma^2 \log = \frac{1}{X} \cdot \sum_{i=1}^X (\text{Log } \bar{Y}_i / \bar{Y})^2$$

where X = the number of individuals in the distribution

\bar{Y}_i = individual income

\bar{Y} = geometric average of the individual incomes;

where the data provide a distribution by classes of income the expression is corrected:

¹¹For more details see Rodolfo Hoffman, "Contribuição à Análise da Distribuição da Renda e da Posse de Terra no Brasil", op.cit.

$$\sigma^2 \log = \frac{1}{X} \cdot \sum_{I=1}^K X_I (\text{Log } \bar{Y}_I / \bar{Y})^2$$

where X = the number of individuals in the distribution;

X_I = the number of individuals in class I ;

\bar{Y}_I = the average income of individuals of class I

\bar{Y} = geometric mean of average incomes of K income classes;

K = the number of income classes.

With reference to the concept of income distribution implicit in $\sigma^2 \log$, it is worth pointing out that this represents a rather static view of the phenomenon insofar as it takes the difference of the squares of the deviations of average incomes in relation to the geometrical average of the distribution. Although this concept is correct and workable it brings nothing new to the understanding of the phenomenon if we wish to treat it in a wider perspective, which at least clarifies further the complex mechanisms underlying the process of income distribution. At first sight, one has the impression that the $\sigma^2 \log$ may be only a mathematical device, although suitably developed which makes more efficient the use of σ^2 , which in Statistics has basic applications.

With respect to the operational characteristics of this measure one may add:

(a) Advantages:

i) easy to interpret since it consists of a measure of dispersion;

ii) disaggregated into components of the variation of income "within" and "between" classes or distributions, in accordance with the property of σ^2 .

- iii) it has the result of minimizing the effect of absolute differences of income through the use of logarithms of the incomes;
 - iv) easy to calculate because it uses the logarithms of the incomes when available.
 - v) it is a measure sensitive to small variations between income distributions.
- (b) Disadvantages:
- i) the non-availability of the logarithms of the incomes makes immediate calculation more difficult;
 - ii) the fact that it is related to the geometric averages of the individual incomes, and not to the arithmetical average of the distribution accentuates dispersal effect of the measure;
 - iii) the non-availability of the individual income data makes the measure less efficient for estimating the degree of inequality of the income distribution.

Further Observations

The summary consideration of Pareto's index, Gini's coefficient of concentration and the variance of logarithms allows one however to highlight some aspects of the evaluation of the inequality of income distributions. These are necessary for the understanding of proper use of the available measures of inequality.

First, that all the measures of inequality of income have an implicit concept of the process of income distribution.¹²

¹²See Anthony B. Atkinson, "On the Measurement of Inequality", op.cit.; O. Elteto and E. Frigyes, "New Income Inequality Measures as Efficient Tools for Causal Analysis and Planning", op.cit.

Secondly, that some measures reflect some aspects of concentration of income in comparison with others. In the cases analysed, Pareto's index tends to be a more efficient measure of inequality for the upper range of the distribution (beginning with the level of average income \bar{Y}_1); whilst Gini's coefficient of concentration and the Variance of Logarithms give a good representation of the inequality of income of the whole distribution.

These two observations indicate that, depending on the characteristics to be brought out in the evaluation of the inequality of an income distribution, there will be appropriate measures to be used.

Finally, it is interesting to point out that, in operational terms, the proper use of a measure of inequality of income requires that it should be at least:

- (a) easy and precise to interpret;
- (b) quickly calculated;
- (c) independent of the form of the distribution;
- (d) sensitive to differences between distributions;
- (e) sensitive to ambiguity between distributions.¹³

¹³ Ambiguity, for example, as in the case of the Lorenz curve when it may happen that the comparison of two of them intersecting provides the same value for the Gini's coefficient of concentration.

CHAPTER II

AN ALTERNATIVE CONCEPT

In this section a further way of visualizing the phenomenon of income distribution is developed in order to propose later a measure of inequality of income. Following the lines of the previous section the concept of income distribution is approached first, followed by the mathematical formulation of this concept, and finally the characteristics of the proposed measure of inequality of income.

The Means of Production and Income Distribution

The classic theories (including Marx) about the value and distribution of goods and services¹⁴ bring out the importance of the ownership of means of production in the determination of the distribution of the incomes, although with varying approaches.¹⁵ In reality, they are practically the only theories (with of course their followers) which try to go beyond the process of distribution of goods and services. They do not take this (the Neo-Classic and Keynesian positions) as being established in a competitive market of factors of production where the prices of goods and services are determined by the marginal productivities of

¹⁴See, especially, Maurice Dobb, Teoria del Valor y de la Distribución desde Adam Smith, op.cit.

¹⁵These variations are clear from Adam Smith to Marx. The latter, in turn, maybe has succeeded best in integrating the theories of production and distribution.

those factors. In the end, they start from socio-economic reality, historically determined; through an analysis of the social structure to distinguish the difference in the ownership of the means of production (social relationships of production) as determinants of the distribution of goods and services produced for the market.

It is evident that political and historical factors exist which are hard to measure - which interacting with the classical economic variables - determine the distribution of individual incomes.¹⁶ However, one may suppose a close relationship between the ownership of the means of production and the structure of political power historically determined. It is reasonable to suppose that both vary in the same direction. The exceptions may become apparent insofar as political power, other things being equal, manifests itself with support of economic power.¹⁷

It is clear that the inter-relation between the ownership of the means of production and the structure of income or rather the historical-structural determination of the latter by the former has an outstanding place in economic writing, particularly in developing countries. The external relationships of these with developed countries clearly reflect an unequal social division of labour in which the latter, with the greater accumulation of capital and technology, determine the distribution of factors in the former and, consequently bring about concentrated income

¹⁶ See Irma Adelman and Cinthia Taft Morris, "¿Quien se beneficia con el desarrollo economico?", Distribución del Ingreso, ed. by Alejandro Foxley (Mexico: Fondo de Cultura Economica, 1974).

¹⁷ This position is reinforced by Jorge Graciarena, "Estrutura de Poder y Distribución del Ingreso en America Latina", ibid.

¹⁸ structures. In short, it may be supposed that the domination or existing influence of the "strong" over the "weak", economically and politically speaking, has been maintained historically, just as much between countries as between their social classes. This inequality of economic opportunities, determined by the marked stratification of the social structure, is reflected objectively in the different appropriations of goods and services produced for the market. It is towards this aspect that the objective of this section is directed which is that of highlighting productive relationships which, within the economic system, determine the relationships of distribution of the factor rewards.

The Question of Interdependence

The ownership of the means of production in a given economic structure determines the distribution of goods and services (real incomes) or of the factor rewards (money incomes). So the allocations of income between individuals (or groups) is a consequence of the entire flow of goods and services produced in a given period. In this way, within the absolute limits of this flow of production, the ownership of the means of production acts and confers on the individuals (or groups) differential access to the rewards. In other words, the rewards of individuals (or groups) are not determined separately in the process of production and therefore differentiated in the process of distribution of goods and services; in reality, they are determined jointly in the productive process in accordance with the possessions of means of production by the individuals (or groups), given the flow of goods and services produced within the period.

¹⁸ See, especially, Celso Furtado, Análise do Modelo Brasileiro (Rio de Janeiro: Civilização Brasileira S.A., 1973).

In the flow of total income generated over a given period, there is an interdependence of incomes of individuals (or of groups). This reflects the fact that mainly for reasons of their respective economic bargaining power (or through the influence of differential ownership of the means of production), the higher allocation of income to some (individuals or groups) necessarily implies a smaller allocation to others.

The Interpretation of the Process of Income Distribution

The previous sections commented on the inter-relationship between the ownership of the means of production and the allocation of income as well as on the question of interdependence of incomes. It becomes now possible to interpret in summary form how the distribution of rewards between individuals (or groups) takes place.

In the first place, the proposition concerning the inter-relationship between the ownership of the means of production, (Z), and the allocation of income, (Y), suggests a direct proportional relationship in the sense that a larger (Z) means a larger (Y), and vice-versa. Without doubt this situation can be generalised although the given historical setting of social-economic structures show specific variations around this proportional relationship. However, it seems impossible to deny that this will always be positive, mainly because of the fact that individuals (or groups) tend to maximize their relative benefits (or rewards) in their economic activities.

¹⁹This aspect is not new, because it is found explicitly in most of the measures of inequality of income. The objective here was to highlight this aspect because of its decisive importance for understanding the process of income distribution - a fact which is generally neglected in other studies.

In the second place, the question of interdependence in the allocations of income between individuals (or groups), (X), tends to show the dynamics of the social relationships of production in the sense that, by the destination of the flow of goods and services produced in a given period, individuals will compete among themselves (or groups) in such a way as to define their respective shares (or allocations). In this case, the ownership of the means of production plays a decisive part since it already reserves for its owners a large part of the rewards obtained in the market. Consequently, the interdependence of incomes will specify the differences in shares of income between individuals (or groups), that is some will necessarily appropriate more than others within the total flow of income generated in the period.

Thirdly, it is necessary to attempt an integrated view of the process of income distribution. The conjunction of the processes of production and distribution²⁰ of goods and services, which we are trying here to make operational, gives the analysis of income distribution a more dynamic perspective with the specific structural implications. In this sense, it starts from the proposition that differential ownership of the means (or factors) of the productive process determines the differential allocation of the rewards of those factors. This difference, in its turn, establishes in the market the distribution of the flow of goods and services produced among the individuals (or groups) through the mechanism of interdependence of incomes, which specify the absolute levels of income differentially allocated.

²⁰The distribution of goods and services in itself only signifies the distribution of real incomes which will be converted into rewards (monetary) by entry to the market.

Thus, the differential allocation of incomes is determined a priori through the differential ownership of the means of production and realised a posteriori through the mechanism of interdependence of incomes.

An Operational Form of the Interpretation of the Process of Income Distribution

Taking the conceptual sequence of the previous sections we present the mathematical formulation of the interpretation of the process of income distribution.

To give effect to that interpretation the following form is chosen:

(a) Let \bar{Z}_I and \bar{Y}_I respectively be the average stock of the means of production and the average income of the group of individuals X_I .

Then, the expression

$$\sum_I X_I \bar{Z}_I \bar{Y}_I, \quad (1)$$

indicates the sum of all groups of individuals in the distribution with the respective allocations of income and means of production.

(b) If all the individuals in the distribution were allocated the same level of income (\bar{Y}), we could write

$$\sum_I X_I \bar{Z}_I \bar{Y}, \quad (2),$$

(c) Verification of the absolute inequality of income between all the individuals in the distribution is possible by utilizing the expressions (1) and (2) as follows:

$$\sum_I X_I \bar{Z}_I \bar{Y}_I - \sum_I X_I \bar{Z}_I \bar{Y} = \sum_I X_I \bar{Z}_I (\bar{Y}_I - \bar{Y}), \quad (3),$$

which final form signifies the weighted sum of all the absolute differences of income of the groups of individuals between the observed distribution and the

ideal distribution,

(d) The relativity²¹ of the measure is obtained by the conjunction of the expressions (1) and (3) or

$$\frac{\sum_I X_I \bar{Z}_I (\bar{Y}_I - \bar{Y})}{\sum_I X_I \bar{Z}_I \bar{Y}_I} \quad (4).$$

Practical Application of the Operation

It is difficult to find operational support for the proposition that the ownership of the means of production determines income distribution since the usual systems of information used do not identify separately the first variable. On the one hand this is because it has become customary to simply reproduce traditional systems of economic information based on Keynesian theory²³ and on the other hand because the very concept of means (or factors) of production makes the question open to different interpretations which seriously impedes its statistical application.

²¹Obtaining positive values for this relative measure depends on the proposition that the ownership of the means of production (stock) is directly proportional to the absolute levels of income of the group of individuals holding them.

²²In reality, the expression (4) can be understood as the weighting of the relative differences of income by the share of the groups of individuals $I = 1, \dots, K$, in the total ownership (stock) (Z) of the means of production, i.e.

$$\frac{\sum_I X_I \bar{Z}_I (\bar{Y}_I - \bar{Y})}{\sum_I X_I \bar{Z}_I \bar{Y}_I} = \frac{\sum_I X_I \bar{Z}_I}{\sum_I X_I \bar{Z}} \cdot \frac{(\bar{Y}_I - \bar{Y})}{\frac{\sum_I X_I \bar{Z}_I}{\sum_I X_I \bar{Z}} \cdot \bar{Y}_I}.$$

²³In this case, Simon Kuznets is one of the exponents of that part which relates to statistical manipulation of macro-economic aggregates.

In this way, except for specific studies previously defined methodologically, the alternative proposed here for measuring the inequality of income is at a disadvantage. The practical application of this alternative is not totally unserviceable because it is possible, with a reasonable degree of confidence, to adapt it suitably to the established conceptual aims.

Thus, starting from the proposition that the ownership of the means of production determines income distribution, and that, consequently, there exists a direct relationship between the two variables (i.e. both vary in the same direction) it becomes possible by a workable method to substitute the first by the second. In other words, one may use the level of income (\bar{Y}_I) of the group of individuals (I) as a "proxy" variable for the average stock of the means of production (\bar{Z}_I) of this group of individuals.²⁴ It is evident that the results will not be the same if the initial

²⁴A more specific interpretation of this proposal could arise from the relations $\bar{Y}_I = r_I \bar{Z}_I$, where r_I is the average rate of return on the stock of the means of production of the group of individuals I.

Given that,

$$\bar{Y}_I = \sum_j r_{Ij} \bar{Z}_{Ij} / \sum_j r_j \bar{Z}_j ,$$

$$I = 1, \dots, K \text{ e } j = 1, \dots, J$$

and,

$$r_j = \sum_{Ij} r_{Ij} X_{Ij} / \sum_j X_{Ij} \text{ for each } j ,$$

then,

$$\bar{Y}_{Ij} = r_j \bar{Z}_{Ij} ,$$

where r_j = the average rate of return on the stock of means of production of the distribution j.

Furthermore, one may adapt the neo-classical variation of the theory of human capital, following the justification above, taking \bar{Z}_I as the average stock of human capital of the group of individuals I and, consequently, altering the significance of r_I and r_j .

expression (4) is used; however it is probable that in periodical comparisons the variations of the two measures (the original and the adapted) follow the same directions.

Thus the indicated operational adaptation reduces to:

$$\frac{\sum_I X_I \bar{Y}_I (\bar{Y}_I - \bar{Y})}{\sum_I X_I \bar{Y}_I^2} = \frac{(\sum_I X_I \bar{Y}_I^2 - \sum_I X_I \bar{Y}_I \bar{Y})}{\sum_I X_I \bar{Y}_I^2}, \quad (5),$$

As $\sum_I X_I \bar{Y}_I = X \bar{Y}$ and substituting in (5) we see

$$\frac{(\sum_I X_I \bar{Y}_I^2 - X \bar{Y}^2)}{\sum_I X_I \bar{Y}_I^2} = \frac{\sum_I X_I (\bar{Y}_I^2 - \bar{Y}^2)}{\sum_I X_I \bar{Y}_I^2}, \quad (6).$$

Given that,

$$\frac{\sum_I X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} = C V^2, \quad (7),$$

the following relationship is obtained,

$$\frac{(\sum_I X_I \bar{Y}_I^2 - \sum_I X_I \bar{Y}^2)}{X \bar{Y}^2} = \frac{\sum_I X_I \bar{Y}_I^2}{XY} - 1 = C V^2$$

and finally,

$$\frac{\sum_I X_I \bar{Y}_I^2}{XY} = C V^2 + 1, \quad (8)$$

Substituting the expressions (7) and (8) in (6), we obtain:

$$\frac{\sum_I X_I (\bar{Y}_I^2 - \bar{Y}^2)}{\sum_I X_I \bar{Y}_I^2} = \frac{\sum_I X_I (\bar{Y}_I^2 - \bar{Y}^2)}{XY} \div \frac{\sum_I X_I \bar{Y}_I^2}{XY} = C V^2 / C V^2 + 1, \quad (9)$$

where $C V^2$ = the square of the coefficient of variation. 25

²⁵Note that the expression $C V^2$ used in the text refers, as was shown, to grouped income data. Also note that $C V^2 = (\sigma/\bar{Y})^2$, where σ = the standard deviation of the distribution.

In this way the expression (9) is the final form of the proposed measure of inequality of incomes.²⁶

The Range of Variation of the Measure

From (9) the limits of variation of the measure are:

(a) Maximum Equality of Income

In this case, where all the individuals in the distribution, whatever their number may be, receive equal incomes, or $\bar{Y}_I = \bar{Y}$, so that $\sigma^2 = 0$, the measure consequently is equal to 0.

(b) Maximum Inequality of Income

In this case, where only 1 (one) individual receives all the income in the distribution, when $X \rightarrow \infty$, $\bar{Y} \rightarrow 0$, and consequently the measure tends to the limit 1.

²⁶In a recent study not known to us till the final stage of this thesis, Jorge de Souza, "Duality and Concentration", Brasília, 1974 (Mimographed), arrived at an alternative measurement of the inequality of income of a distribution, whose final expression is the same as the one suggested here. Though the results of both works may be the same in their final expression, the two attempts are not exclusive since the concepts and methodologies used by the two authors are themselves different. While Souza starts with the property of the "duality of function", econometric approach, our work takes an economic approach. Far from constituting similar works, the two attempts are on the one hand specific, to the extent to which the fields of analysis are different, and on the other hand complementary to the extent to which the statistical/mathematical framework is necessary in giving operational form to the measurable economic interpretations. Finally, it is relevant to point out, throughout our work, that our concern was directed more to the concept of the phenomenon of income distribution, evidently capable of being made operational than to elaborating a measure of inequality as such.

$$\text{In brief, } 0 \leq \frac{CV^2}{CV^2 + 1} < 1.$$

Graphical Presentation

The expression (6) permits the following graphical presentation:

$$\text{from } \frac{CV^2}{CV^2 + 1} = \frac{\sum_{I=1}^K X_I \bar{Y}_I^2 - X \bar{Y}^2}{\sum_{I=1}^K X_I \bar{Y}_I^2} = \frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{\sum_{I=1}^K X_I \bar{Y}_I^2}$$

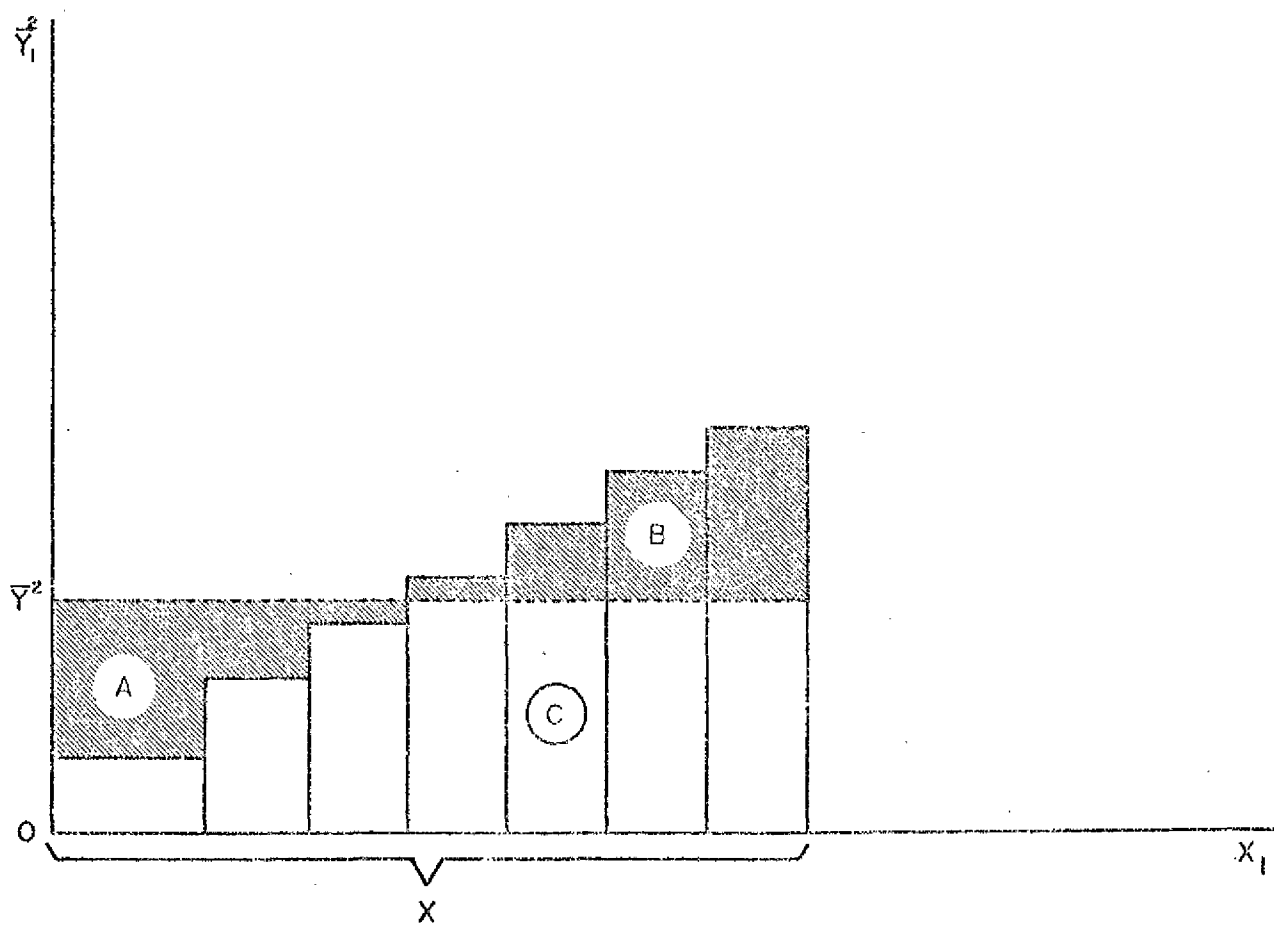
we have (see figure 3).

$$\begin{aligned} \text{Thus, } \frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{\sum_{I=1}^K X_I \bar{Y}_I^2} &= \frac{(\text{AREA C} + \text{AREA B}) - (\text{AREA C} + \text{AREA A})}{\text{AREA C} + \text{AREA B}} \\ &= \frac{\text{AREA B} + \text{AREA A}}{\text{AREA C} + \text{AREA B}}, \quad \text{where AREA A} = |-\text{AREA A}| \end{aligned}$$

Disaggregation of the Measure

An important characteristic of the proposed measure of inequality of income is the fact of its being disaggregated into two components, one which measures the inequality of income observed "within" the classes or distributions and the other which measures the inequality of income "between" the classes or distributions. In the case of measuring the inequality of income "within" and "between" the classes of a distribution, it is indispensable to have data on individual incomes, which will furnish the necessary information for the disaggregation.

FIGURE 3

GRAPHICAL PRESENTATION OF $CV^2/CV^2 + 1$ 

The idea underlying the disaggregation is that the degree of inequality of income of a given distribution is formed by the component of inequality between the incomes of the individuals included "within" each class I and by the component of inequality "between" incomes of individuals in each class I (in this case, the incomes of individuals of each class I are represented by the average income (\bar{Y}_I) of the respective class). In the case where the distribution is formed by J distributions, the measure of inequality of income of this distribution will be also disaggregated into the two previous components. One which would measure the inequality of income between the incomes of the population classified "within" each distribution j and the other which would measure the inequality "between" the incomes of the populations of the j distributions (in this case, the incomes of the populations of each distribution j would be represented by the average income (\bar{Y}_j) of the respective distribution). So summing up the proposed measure of inequality of income can be explained just as much by the components of inequality between the individual incomes "within" and "between" each class I - when individual income data are available - , as by the components of inequality between the individual incomes "within" and "between" each distribution j - when the distribution of total income is formed by J distributions.

This characteristic of the proposed measure of inequality of income is of great value in the sense that it gives greater explanatory power to the measure and provides additional elements to the analysis of the inequality of income.

A practical example of the disaggregation can be imagined when one wishes to evaluate the degree of inequality of income of a given economy, and one has available, at the same time, information about the income distributions of economic sectors. In this way, the measure of inequality of

income calculated for this economy will be formed by the component of inequality of income "between" the sectors and by the component of inequality of income "within" the sectors. The same example applies to the case where one also has information about the income distribution of the social classes in this economy. In this case, the same value of the measure of inequality of income of the economy will be formed by the component of inequality of income "within" each social class and by the component of inequality of income "between" the social classes.

This being so, given an aggregate income distribution, formed by K classes or J distributions, one can disaggregate the proposed measure of inequality of income as follows:

(1) By Income Classes

Taking the expression of the measure in (9), and transforming it for individual income data, we have

$$\frac{C V_T^2}{C V_T^2 + 1} = \frac{\sigma^2 / \bar{Y}^2}{\sigma^2 / \bar{Y}^2 + 1} = \frac{\sum_{i=1}^X (\bar{Y}_i^2 - \bar{Y}^2) / X \bar{Y}^2}{\sum_{i=1}^X (\bar{Y}_i^2 - \bar{Y}^2) / X \bar{Y}^2 + 1}$$

taking solely the numerator we have

$$\frac{\sum_{i=1}^X (\bar{Y}_i^2 - \bar{Y}^2)}{X \bar{Y}^2} = \frac{\sum_{I=1}^K X_I (\bar{Y}_I - \bar{Y})^2}{X \bar{Y}^2} + \frac{\sum_{I=1}^K X_I \bar{Y}_I^2}{X \bar{Y}^2} - \frac{\sum_{i=1}^I (\bar{Y}_{iI}^2 - \bar{Y}_I^2)}{X_I \bar{Y}_I^2} \quad 27$$

27

A proof of this disaggregation is given in the appendix.

substituting the disaggregated numerator in the measure, we have

$$\frac{C V_T^2}{C V_T^2 + 1} = \frac{\frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} + \frac{\sum_{I=1}^K X_I \bar{Y}_I^2}{X \bar{Y}^2} \cdot \frac{\sum_{i=1}^I (\bar{Y}_{iI}^2 - \bar{Y}_I^2)}{X_I \bar{Y}_I^2}}{\frac{\sum_{i=1}^I (\bar{Y}_i^2 - \bar{Y}^2)}{X \bar{Y}^2} + 1}$$

$$\text{or } \frac{C V_T^2}{C V_T^2 + 1} = \frac{C V_B^2 + \frac{\sum_{I=1}^K X_I \bar{Y}_I^2}{X \bar{Y}^2} \cdot C V_W^2}{C V_T^2 + 1};$$

where,

$$C V_B^2 = \frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} = \text{component of income inequality "between" the classes of the distribution;}$$

X_I = total of individuals in income class I ;

X = total of individuals in the distribution;

\bar{Y}_I = average income of individuals in class I ;

\bar{Y} = average income of the distribution;

K = total of income classes.

$$C V_W^2 = \frac{\sum_{i=1}^I (\bar{Y}_{iI}^2 - \bar{Y}_I^2)}{X_I \bar{Y}_I^2} = \text{component of income inequality "within" the classes of the distribution;}$$

X_I = total of individuals in the income class I ,
 X = total of individuals in the distribution;
 \bar{Y}_{iI} = individuals income of class I ;
 \bar{Y}_I = average income of the individuals in class I ;
 K = number of income classes.

$$C V_T^2 = \frac{\sum_{i=1}^K X_I (\bar{Y}_{iI}^2 - \bar{Y}_I^2)}{X \bar{Y}^2} = \text{component of income inequality of the distribution,}$$

\bar{Y}_i = individual income;
 \bar{Y} = average income of the distribution;
 X = total number of individuals in the distribution.

(2) By Distributions

In this case, the disaggregation of the measure is made by starting from a distribution built up by the combination of J distributions. Thus, taking only the numerator from the expression (9), we have

$$\frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} = \frac{\sum_{j=1}^J X_j (\bar{Y}_j^2 - \bar{Y}^2)}{X \bar{Y}^2} + \frac{\sum_{j=1}^J X_j \bar{Y}_j^2}{X \bar{Y}^2} \cdot \frac{\sum_{I=1}^{K_j} X_{Ij} (\bar{Y}_{Ij}^2 - \bar{Y}_j^2)}{X_j \bar{Y}_j^2} \quad 28$$

substituting the disaggregated numerator in the measure we have

$$C V_T^2 = \frac{\sum_{j=1}^J X_j (\bar{Y}_j^2 - \bar{Y}^2)}{X \bar{Y}^2} + \frac{\sum_{j=1}^J X_j \bar{Y}_j^2}{X \bar{Y}^2} \cdot \frac{\sum_{I=1}^{K_j} X_{Ij} (\bar{Y}_{Ij}^2 - \bar{Y}_j^2)}{X_j \bar{Y}_j^2}$$

$$C V_T^2 + 1 = \frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} + 1$$

²⁸The proof of this disaggregation follows the logic which was used in the appendix.

$$\text{or } \frac{C V_T^2}{C V_T^2 + 1} = \frac{C V_B^2 + \frac{\sum_{j=1}^J X_j \bar{Y}_j^2}{X \bar{Y}^2} (C V_W^2)}{C V_T^2 + 1}$$

where,

$$C V_B^2 = \frac{\sum_{j=1}^J X_j (\bar{Y}_j^2 - \bar{Y}^2)}{X \bar{Y}^2} = \text{component of income inequality "between" the } \underline{J} \text{ distributions;}$$

X_j = total number of individuals in the distribution \underline{j} ;

X = number of individuals of the \underline{J} distributions;

\bar{Y}_j = average income of distribution \underline{j} ;

\bar{Y} = average income of the \underline{J} distributions;

J = number of distributions which together form the total distribution;

$$C V_W^2 = \frac{\sum_{I=1}^{K_j} X_{Ij} (\bar{Y}_{Ij}^2 - \bar{Y}_j^2)}{X_j \bar{Y}_j^2} = \text{component of income inequality "within" the } \underline{J} \text{ distributions;}$$

X_j = total number of individuals of the distribution \underline{j} ;

X = total number of individuals of the \underline{J} distributions;

\bar{Y}_j = average income of the distribution \underline{j} ;

\bar{Y} = average income of the \underline{J} distributions;

\bar{Y}_{Ij} = average income of class I of distribution \underline{j} ;

K_j = total of income classes of the distribution \underline{j} ;

$$C V_T^2 = \frac{\sum_{I=1}^K X_I (\bar{Y}_I^2 - \bar{Y}^2)}{X \bar{Y}^2} = \text{component of income inequality of the aggregated distribution.}$$

- X_I = number of individuals in class I of the aggregated distribution;
 X = number of individuals in the aggregated distribution;
 \bar{Y}_I = average income of class I;
 \bar{Y} = average income of the aggregated distribution;
 K = total number of income classes.

Properties of the Measure

We try next to comment on the properties of the proposed measure of inequality of income, given the observations made in the previous sections, as a final procedure to appreciate and evaluate its ability to estimate the concentration of income in a given distribution. These properties are arranged in two groups, one which brings together its characteristics²⁹ considered advantageous, and the other characteristics considered disadvantageous, or at least as limiting its analytical power.

(a) Advantages

i) The measure has a structural concept of the process of income distribution to the extent to which it adopts the approach of classical economic theory.

This concept allows a wider interpretation of the working of the processes of generation and income distribution in a given economic system.

ii) The measure is decomposed in components of inequality of income "within" classes or distributions and "between" classes or distributions.

²⁹These characteristics are grouped on the basis of the analysis of similar measures as well as a critical appreciation of other authors concerning income distribution.

This property confers on the measure greater power to analyse and interpret the process of income distribution.

iii) The measure is independent of the form taken by the distribution.

This property is due to the fact that it is not necessary to know the form of the distribution in order to apply the measure.

iv) The measure does not show ambiguities of interpretation of degrees of inequality of income when two or more distributions are compared.

This property is acquired through the characteristics of CV^2 , according to which the variance and the average will only be equal in the case of having two or more similar income distributions, that is to say identical or proportional.

v) The measure is sensitive to small differences between income distributions.

This property arises from the characteristics of the squares of the weighted differences of income.

vi) The simplicity of its meaning: the ratio between the sum of the squares of the weighted differences of income (of individuals or groups of individuals and the average income of the distribution) and the sum of the squares of the weighted incomes (of individuals or groups of individuals).

vii) The measure shows well the inequality of income observed in all the distribution.

viii) The method of calculation of the measure is simple and quick.

(b) Disadvantages

i) The measure only achieves greater precision in estimating the degree of inequality of income of a distribution when individuals' income data are available.

This is explained by the fact that the measure takes the simplified form of the ratio between two factors in which CV^2 occurs. This, being a measure of dispersion, inevitably requires all the data about the distribution in order to make it a precise measuring instrument.

ii) In the disaggregation of the measure between components of inequality of income "between" and "within", the sum of the weights is at the minimum equal to 1.

This fact means that this sum attains the value $1 + C V_E^2$ ³⁰, which implies that the component of inequality of income "within", including its weights, depends on the value taken by the component of inequality of income "between".

³⁰Henry Theil, Economics and Information Theory (Amsterdam: North-Holland, 1967), p. 125.

CHAPTER III

APPLICATION OF THE MEASURE

In this section we make use of the proposed measure to estimate the relative inequality of income in two cases: (1) employing the overall measure and (2) employing its property of disaggregation. As the objective is only to test the power of analysis and interpretation of the measure in two specific cases, we shall not try to enter into the details about the implications of the results obtained.

(1) We shall now use the proposed measure to estimate the degree of concentration existing in the income distributions of Brazil from 1960 to 1970:

(a) The income distribution for 1960 was obtained in: Brazilian Institute of Geography and Statistics, Demographic Census - Preliminary Results - Special Series, 1960 (Rio de Janeiro: IBGE Printing Office, 1965), p.9. The mid-points of the income classes with the exception of the last were used. For calculating the average income of the latter a Pareto curve was fitted to the last two ranges of income, giving the coefficient $b = 1,8764$.

(b) The income distribution for 1970 was obtained in: Brazilian Institute of Geography and Statistics, Advanced Tabulations of the Demographic Census, 1970 (Rio de Janeiro: IBGE Printing Office, 1973), p.6. As in the previous distribution, the mid-points of the ranges of income below the last were used. For calculating

the average income of this a Pareto curve was fitted to the last two ranges of income, giving the coefficient $b = 1,6613$.

The two income distributions with the respective values of the measure are shown below in tables 1 and 2.

As will be seen, the indicator of concentration of income in Brazil in the decade of the 60s through the estimate given by the proposed measure, agrees with the results of other works. Thus for the years 1960 and 1970 were found respectively: Fishlow (0.59 and 0.63), Langoni (0.50 and 0.56), Duarte and Hoffman (0.49 and 0.57) and Kingston and Kingston (0.42 and 0.53) - values corresponding to the Gini's coefficient of concentration.

Of course the different values obtained reflect not only the different estimates of average incomes of the classes, but also adaptations in the content of the data used (inclusion or not of the individuals who do not receive incomes). However, it is not the objective of this work to make qualitative comparisons of the different estimates of income concentration. It is only intended to show the appropriateness of the proposed measure, as opposed to instruments traditionally used.

(2) Finally, we used the proposed measure to estimate the degree of concentration existing in the distributions by salary groups of the formal urban sector for regions of Brazil in the period 1970/73.

The distributions by salary groups were obtained from the 2/3 Law, National Manpower Department, Ministry of Labour. This information refers to the

TABLE I
 DISTRIBUTION OF INDIVIDUAL INCOMES^a
 IN BRAZIL - 1960
 (In Current Cruzeiros)

| Including Individuals Without Incomes | | | Excluding Individuals Without Incomes | | |
|--|--------------------------------------|----------|--|--------------------------------------|----------|
| Ranges of Income | Average Income (\bar{Y}_I) | b EAP | Ranges of Income | Average Income (\bar{Y}_I) | b EAP |
| 0 | 0 | 3 322 | 2,1 | 1,05 | 4 788 |
| 2,1 | 1,05 | 4 788 | 2,1 - 3,3 | 2,70 | 3 242 |
| 2,1 - 3,3 | 2,70 | 3 242 | 3,3 - 4,5 | 3,90 | 2 464 |
| 3,3 - 4,5 | 3,90 | 2 464 | 4,5 - 6,0 | 5,25 | 2 887 |
| 4,5 - 6,0 | 5,25 | 2 887 | 6,0 - 10,0 | 8,00 | 3 180 |
| 6,0 - 10,0 | 8,00 | 3 180 | 10,0 - 20,0 | 15,00 | 1 733 |
| 10,0 - 20,0 | 15,00 | 1 733 | 20,0 - + | 42,82 | 649 |
| 20,0 - + | 42,82 | 649 | | | |
| $CV^2/CV^2 + 1$ | 0,677 | | $CV^2/CV^2 + 1$ | 0,620 | |

^aSource: Brazilian Institute of Geography and Statistics.

^bEconomically active population in thousands.

TABLE 2
 DISTRIBUTION OF INDIVIDUAL INCOMES^a
 IN BRAZIL - 1970
 (In Current Cruzeiros)

| Including Individuals Without Incomes | | | Excluding Individuals Without Incomes | | |
|--|---------------------------------------|---------------------|--|---------------------------------------|---------------------|
| Ranges of Income | Average Incomes (\bar{Y}_I) | ^b EAP | Ranges of Income | Average Incomes (\bar{Y}_I) | ^b EAP |
| 0 | 0 | 2 755 | 100 | 50 | 9 374 |
| 100 | 50 | 9 374 | 101 - 150 | 125 | 3 770 |
| 101 - 150 | 125 | 3 770 | 151 - 200 | 175 | 4 604 |
| 151 - 200 | 175 | 4 604 | 201 - 250 | 225 | 1 321 |
| 201 - 250 | 225 | 1 321 | 251 - 500 | 375 | 4 307 |
| 251 - 500 | 375 | 4 307 | 501 - 1 000 | 750 | 1 738 |
| 501 - 1 000 | 750 | 1 738 | 1 001 - 2 000 | 1 500 | 660 |
| 1 001 - 2 000 | 1 500 | 660 | 2 001 - + | 5 027 | 306 |
| 2 001 - + | 5 027 | 306 | | | |
| $CV^2/CV^2 + 1$ | 0,824 | | $CV^2/CV^2 + 1$ | 0,806 | |

^aSource: Brazilian Institute of Geography and Statistics.

^bEconomically active population in thousands.

wages which enterprises declared as paid to their employees and is usually supplied in April each year, by means of standard questionnaires which by law the enterprises must send to the Regional Labour Delegations.

These data refer to the "formal" sector of the urban economy, i.e. they cover industry, commerce and services (excluding unpaid, personal and own-account services). In terms of representativeness the data covered about 45% of the individuals declared to be employed in the population census of 1970.

Moreover, the 2/3 Law does not furnish the remuneration of employers and other administrative and technical professions (consultants, advisers, administrators, high-level and similar technicians), liberal professions and other workers on their own account, nor individuals connected with "informal" activities (those exercised outside the legal framework which regulates the urban economy).

We present in Table 3 and the corresponding figure the values shown by the measure with the respective components of inequality of total income, within and between distributions. The distributions of wages by groups and regions are shown in Tables 4 and 7.

The interpretation of the data contained in Table 3 and the corresponding figure suggests the following observations:

(a) The tendency to concentration of income occurring in the period 1970/73 was increasing slightly - 0.602 in 1970 to 0.606 in 1973 -, although with a very low growth - just under 1%. Although this result for Brazil may on the one hand show a very slow growth of concentration of wages, it also shows a certain rigidity in the wage structure in

TABLE
DISAGGREGATION C

| Component | Year | 1970 | |
|-----------|-------------|-------|-------|
| | | Value | % |
| Total | | 0,502 | 100,0 |
| Between | | 0,003 | 1,4 |
| Within | | 0,594 | 98,6 |
| Regions: | North | 0,005 | 0,9 |
| | Northeast | 0,043 | 7,2 |
| | Southeast | 0,494 | 81,9 |
| | South | 0,044 | 7,3 |
| | Center West | 0,008 | 1,3 |

^a Calculated from tables 4 to 7.

3
 DF $CV^2/CV^2 + 1^a$

| 1971 | | 1972 | | 1973 | |
|-------|-------|-------|-------|-------|-------|
| Value | % | Value | % | Value | % |
| 0,601 | 100,0 | 0,605 | 100,0 | 0,606 | 100,0 |
| 0,008 | 1,3 | 0,010 | 1,6 | 0,009 | 1,5 |
| 0,593 | 98,7 | 0,595 | 98,4 | 0,597 | 98,5 |
| 0,009 | 1,5 | 0,007 | 1,1 | 0,006 | 1,0 |
| 0,041 | 6,8 | 0,035 | 5,8 | 0,038 | 6,3 |
| 0,485 | 81,1 | 0,497 | 82,2 | 0,493 | 81,7 |
| 0,052 | 8,7 | 0,046 | 7,7 | 0,046 | 7,6 |
| 0,006 | 1,0 | 0,010 | 1,6 | 0,014 | 2,3 |

FIGURE 4
DISAGGREGATION OF $CV^2 / CV^2 + 1$

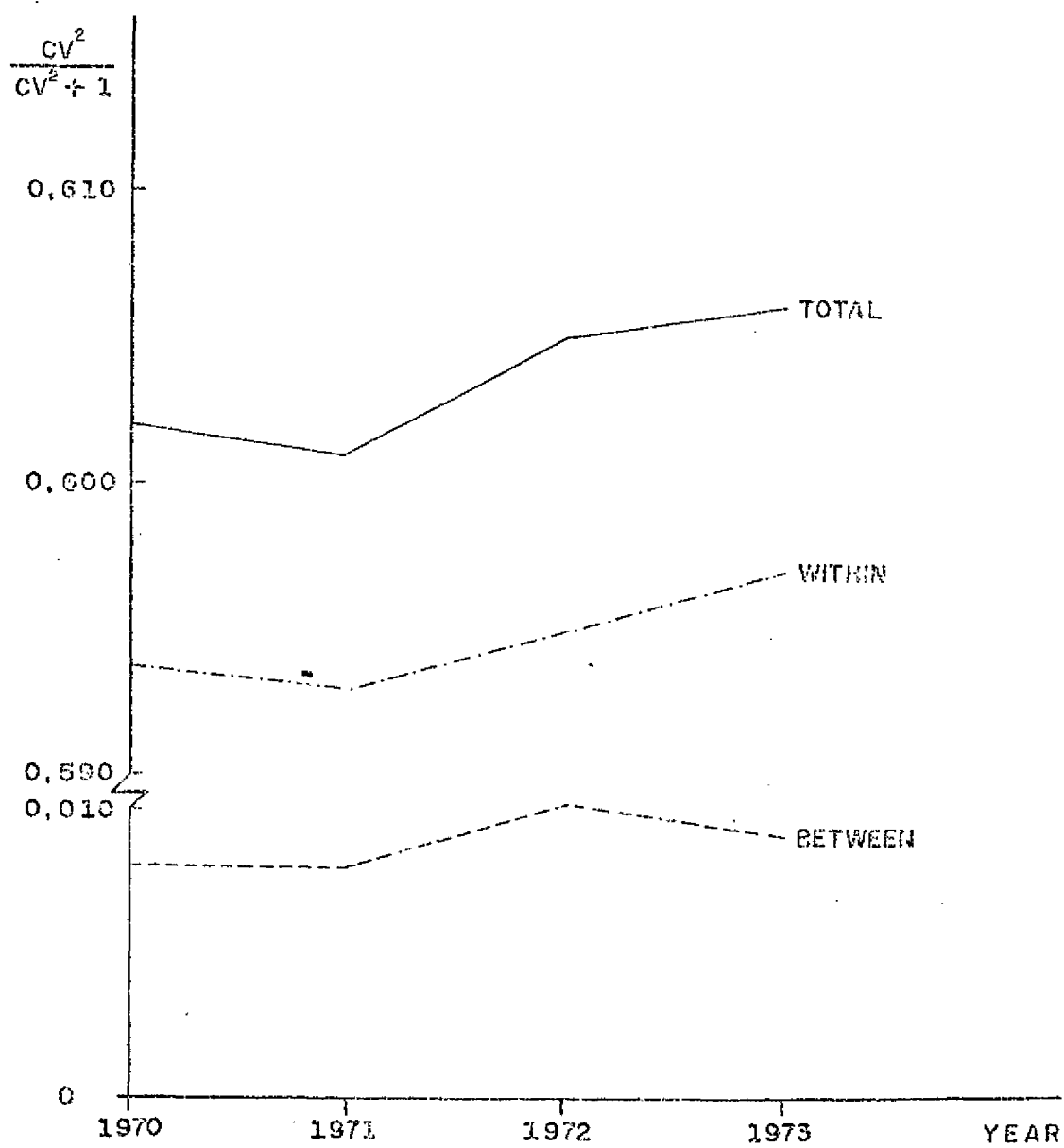


TABLE 4
DISTRIBUTION OF WAGES^a IN THE FORMAL URBAN SECTOR^b
BY REGIONS - 1970

| Ranges of Wages ^c | Regions | | | | | | | | | | Brazil | |
|------------------------------|-------------|--------|-------------|---------|-------------|-----------|-------------|---------|-------------|---------|-------------|-----------|
| | North | | Northeast | | Southeast | | South | | Center West | | \bar{Y}_I | X_I |
| | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I | | |
| -140 | 121,4 | 24 280 | 113,4 | 230 287 | 107,0 | 193 033 | 116,8 | 91 990 | 119,2 | 30 662 | 112,4 | 570 252 |
| 140-159 | 148,0 | 11 425 | 148,6 | 54 856 | 152,8 | 542 593 | 147,1 | 206 883 | 149,3 | 26 290 | 151,0 | 842 047 |
| 160-179 | 165,3 | 6 123 | 167,7 | 32 233 | 168,3 | 372 868 | 168,7 | 89 538 | 168,8 | 12 594 | 168,3 | 513 356 |
| 180-199 | 189,9 | 3 611 | 188,4 | 26 210 | 188,4 | 272 728 | 187,9 | 79 955 | 187,9 | 8 428 | 188,3 | 390 932 |
| 200-299 | 244,0 | 12 272 | 240,6 | 90 282 | 240,6 | 880 451 | 239,3 | 189 675 | 237,5 | 33 143 | 240,3 | 1 205 823 |
| 300-399 | 341,8 | 6 030 | 399,4 | 45 097 | 343,8 | 567 047 | 338,9 | 85 626 | 339,8 | 15 894 | 342,8 | 719 694 |
| 400-499 | 442,4 | 3 582 | 442,6 | 25 772 | 445,5 | 331 817 | 441,3 | 44 473 | 427,0 | 9 429 | 444,4 | 415 073 |
| 500-599 | 545,1 | 2 646 | 541,8 | 16 132 | 542,9 | 219 295 | 540,0 | 27 943 | 537,5 | 4 593 | 542,5 | 270 608 |
| 600-699 | 648,3 | 1 614 | 639,9 | 11 072 | 642,2 | 143 706 | 643,2 | 20 118 | 649,4 | 4 500 | 642,4 | 181 010 |
| 700-799 | 739,9 | 877 | 746,6 | 8 451 | 742,3 | 98 242 | 743,0 | 11 235 | 747,4 | 1 853 | 742,7 | 120 658 |
| 800-899 | 844,8 | 1 270 | 843,3 | 7 295 | 842,4 | 71 984 | 839,2 | 8 256 | 842,7 | 1 190 | 842,2 | 89 995 |
| 900-999 | 950,9 | 530 | 946,9 | 4 521 | 945,3 | 47 372 | 946,8 | 6 031 | 940,8 | 795 | 945,6 | 59 249 |
| 1000-1199 | 1 078,1 | 1 076 | 1 088,3 | 10 068 | 1 085,0 | 76 902 | 1 086,7 | 9 363 | 1 078,7 | 1 840 | 1 085,3 | 99 249 |
| 1200-1599 | 1 392,3 | 1 206 | 1 374,8 | 10 361 | 1 371,4 | 81 376 | 1 364,4 | 9 875 | 1 378,5 | 2 292 | 1 371,4 | 105 110 |
| 1600-1999 | 1 779,8 | 653 | 1 771,9 | 4 591 | 1 766,3 | 42 579 | 1 758,4 | 4 695 | 1 751,1 | 1 049 | 1 766,0 | 53 567 |
| 2000-2399 | 2 195,3 | 229 | 2 162,5 | 2 328 | 2 169,0 | 25 383 | 2 158,0 | 2 512 | 2 071,2 | 492 | 2 166,3 | 30 944 |
| 2400 and + | 3 460,5 | 469 | 3 314,8 | 4 079 | 3 807,1 | 53 761 | 3 500,4 | 3 416 | 3 165,2 | 616 | 3 749,1 | 62 341 |
| $\bar{Y}, \Sigma X_I$ | 305,6 | 77 893 | 304,1 | 583 634 | 402,8 | 4 021 137 | 295,7 | 891 584 | 302,4 | 155 660 | 384,6 | 5 729 908 |
| $cv^2/cv^2 + 1$ | 0,611 | | 0,631 | | 0,597 | | 0,546 | | 0,540 | | 0,602 | |

^a Source: Ministry of Labor - National Department of Labor - 2/3 Law.

^b Industry, commerce and services (excluding unpaid, personal and own - account services).

^c Wages declared to be paid by enterprises to their employees (Cr\$ 1,00).

DISTRIBUTION OF WAGES
BY

| Ranges of Wages ^c | North | | Northeast | | South |
|------------------------------------|-------------|---------|-------------|---------|-------------|
| | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I |
| | -140 | 128,3 | 5 363 | 129,1 | 134 437 |
| 140-159 | 146,7 | 20 859 | 148,2 | 91 027 | 149,4 |
| 160-179 | 167,3 | 9 654 | 167,7 | 46 681 | 175,3 |
| 180-199 | 189,4 | 11 305 | 188,9 | 38 244 | 189,0 |
| 200-299 | 240,9 | 17 716 | 241,1 | 101 952 | 238,1 |
| 300-399 | 343,3 | 11 657 | 340,7 | 61 852 | 342,7 |
| 400-499 | 443,7 | 7 171 | 442,5 | 33 877 | 443,8 |
| 500-599 | 545,4 | 4 049 | 539,6 | 20 103 | 542,9 |
| 600-699 | 646,6 | 2 361 | 642,9 | 15 351 | 641,4 |
| 700-799 | 748,3 | 2 085 | 744,8 | 12 156 | 743,0 |
| 800-899 | 848,7 | 1 489 | 843,4 | 9 529 | 842,3 |
| 900-999 | 943,6 | 1 304 | 944,3 | 7 988 | 945,3 |
| 1000-1199 | 1 072,3 | 2 441 | 1 085,0 | 12 764 | 1 082,0 |
| 1200-1599 | 1 371,1 | 2 013 | 1 377,4 | 11 722 | 1 369,3 |
| 1600-1999 | 1 756,7 | 1 037 | 1 778,2 | 6 317 | 1 777,3 |
| 2000-2399 | 2 167,2 | 823 | 2 163,3 | 4 255 | 2 170,9 |
| 2400 and + | 3 971,8 | 1 209 | 3 394,0 | 6 556 | 3 936,9 |
| $\bar{Y}, \Sigma X_I$ | 395,3 | 102 536 | 365,3 | 614 811 | 500,4 |
| CV^2/CV^2+1 | 0,630 | | 0,613 | | 0, |

^aSource: Ministry of Labor - National Department of

^b Industry, commerce and services (excluding unpa

^c Wages declared to be paid by enterprises to the

TABLE 5

a IN THE FORMAL URBAN SECTOR^b
REGIONS - 1971

| Regions | | | | | Brazil | |
|-----------|-------------|---------|-------------|---------|-------------|-----------|
| East | South | | Center West | | | |
| X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I |
| 116 560 | 111,8 | 36 115 | 111,3 | 2 542 | 118,5 | 295 017 |
| 91 362 | 153,9 | 51 954 | 148,5 | 17 362 | 149,6 | 272 564 |
| 182 879 | 170,2 | 141 022 | 176,1 | 17 777 | 172,4 | 398 013 |
| 510 164 | 186,7 | 134 170 | 189,1 | 9 572 | 188,6 | 703 455 |
| 1 139 676 | 238,1 | 283 721 | 241,6 | 33 068 | 238,4 | 1 576 133 |
| 600 596 | 339,8 | 123 707 | 337,1 | 17 191 | 342,0 | 815 003 |
| 426 931 | 439,9 | 65 777 | 428,8 | 12 310 | 442,9 | 546 066 |
| 248 194 | 541,0 | 37 356 | 541,4 | 2 681 | 542,5 | 312 383 |
| 186 780 | 640,9 | 23 274 | 654,3 | 4 538 | 641,8 | 232 304 |
| 134 467 | 739,7 | 16 597 | 742,4 | 782 | 742,9 | 166 087 |
| 100 975 | 834,5 | 13 799 | 852,7 | 533 | 841,7 | 126 325 |
| 70 671 | 943,3 | 9 483 | 944,2 | 1 091 | 945,0 | 90 537 |
| 95 443 | 1 078,1 | 16 338 | 1 066,5 | 945 | 1 081,5 | 127 931 |
| 107 679 | 1 363,3 | 16 864 | 1 383,9 | 1 314 | 1 369,4 | 139 592 |
| 58 071 | 1 764,4 | 8 895 | 1 797,2 | 934 | 1 775,8 | 75 254 |
| 36 080 | 2 167,2 | 4 833 | 2 173,0 | 420 | 2 169,8 | 46 411 |
| 85 853 | 3 556,5 | 7 509 | 3 243,0 | 1 251 | 3 866,9 | 102 378 |
| 4 192 381 | 364,2 | 991 414 | 344,5 | 124 311 | 459,2 | 6 025 453 |
| 597 | 0,558 | | 0,563 | | 0,601 | |

of Labor - 2/3 Law.

id, personal and own - account services).

ir employees (Cr\$ 1,00).

DISTRIBUTION OF WAGES
BY

| Ranges of Wages ^c | North | | Northeast | | South |
|------------------------------------|-------------|---------|-------------|---------|-------------|
| | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I |
| | -140 | 112,5 | 646 | 108,0 | 11 157 |
| 140-159 | 145,5 | 430 | 151,0 | 47 091 | 148,6 |
| 160-179 | 172,3 | 5 052 | 166,7 | 136 392 | 168,8 |
| 180-199 | 187,8 | 24 077 | 187,8 | 63 411 | 187,2 |
| 200-299 | 237,2 | 33 084 | 239,5 | 139 493 | 244,9 |
| 300-399 | 343,7 | 16 142 | 339,9 | 75 149 | 341,1 |
| 400-499 | 446,6 | 10 518 | 439,7 | 47 956 | 444,6 |
| 500-599 | 541,3 | 5 570 | 540,1 | 29 222 | 543,5 |
| 600-699 | 642,7 | 4 217 | 640,3 | 19 485 | 642,0 |
| 700-799 | 739,9 | 2 481 | 743,1 | 12 953 | 743,9 |
| 800-899 | 843,9 | 1 934 | 845,3 | 10 717 | 843,9 |
| 900-999 | 947,3 | 1 880 | 944,7 | 9 694 | 944,9 |
| 1000-1199 | 1 087,9 | 2 051 | 1 090,3 | 14 324 | 1 086,4 |
| 1200-1599 | 1 351,1 | 3 265 | 1 351,6 | 16 450 | 1 368,7 |
| 1600-1999 | 1 771,8 | 1 387 | 1 764,1 | 8 755 | 1 771,2 |
| 2000-2399 | 2 187,3 | 1 061 | 2 166,5 | 5 345 | 2 172,7 |
| 2400 and + | 3 503,9 | 2 778 | 3 684,8 | 12 007 | 4 223,7 |
| $\bar{Y}, \Sigma X_I$ | 483,8 | 117 073 | 437,2 | 659 601 | 646,8 |
| $CV^2/CV^2 + 1$ | 0,598 | | 0,625 | | 0 |

^aSource: Ministry of Labor - National Department of
^b Industry, commerce and services (excluding unpaid
^c Wages declared to be paid by enterprises to their

TABLE 6

^a IN THE FORMAL URBAN SECTOR^b
REGIONS - 1972

| Regions | | | | | Brazil | |
|----------------|-------------|----------------|-------------|----------------|-------------|----------------|
| east | South | | Center West | | | |
| X _I | \bar{Y}_I | X _I | \bar{Y}_I | X _I | \bar{Y}_I | X _I |
| 90 101 | 111,6 | 25 234 | 113,5 | 2 623 | 115,6 | 129 761 |
| 29 323 | 151,3 | 15 222 | 148,7 | 1 301 | 150,2 | 93 367 |
| 59 188 | 169,0 | 13 451 | 172,5 | 22 525 | 168,0 | 236 608 |
| 61 615 | 187,8 | 52 886 | 186,8 | 15 588 | 187,6 | 217 577 |
| 1 488 859 | 236,7 | 491 510 | 235,7 | 65 546 | 242,4 | 2 218 492 |
| 763 089 | 339,3 | 183 075 | 339,6 | 32 702 | 340,7 | 1 070 157 |
| 512 943 | 440,6 | 99 044 | 445,3 | 24 502 | 443,7 | 694 963 |
| 367 840 | 539,4 | 58 895 | 536,9 | 9 262 | 542,6 | 470 789 |
| 256 579 | 640,4 | 39 423 | 645,3 | 8 739 | 641,8 | 328 443 |
| 203 060 | 739,2 | 27 672 | 741,2 | 4 693 | 743,3 | 250 859 |
| 146 435 | 839,5 | 19 345 | 841,4 | 3 193 | 843,5 | 181 624 |
| 118 970 | 941,8 | 15 673 | 937,8 | 1 829 | 944,5 | 148 046 |
| 167 590 | 1 085,5 | 21 677 | 1 079,3 | 3 276 | 1 086,5 | 208 918 |
| 173 850 | 1 360,2 | 23 281 | 1 371,4 | 4 774 | 1 367,1 | 221 620 |
| 97 422 | 1 768,1 | 13 466 | 1 755,5 | 2 235 | 1 770,1 | 123 765 |
| 62 248 | 2 151,9 | 7 437 | 2 131,2 | 1 823 | 2 169,5 | 77 914 |
| 169 349 | 3 750,1 | 14 853 | 3 460,9 | 3 846 | 4 132,8 | 202 833 |
| 4 768 461 | 444,3 | 1122144 | 453,1 | 208 457 | 585,0 | 6 875 736 |
| ,598 | 0,558 | | 0,573 | | 0,605 | |

f Labor - 2/3 Law.

d, personal and own - account services).

r employees (Cr\$ 1,00).

TABLE
DISTRIBUTION OF WAGES^a IN THE
BY REGIONS

| Ranges of Wages ^c | Regions | | | | | |
|------------------------------------|-------------|---------|-------------|---------|-------------|-----------|
| | North | | Northeast | | Southeast | |
| | \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I |
| -140 | 94,7 | 1 087 | 99,9 | 10 397 | 69,6 | 66 097 |
| 140-159 | 148,8 | 265 | 149,1 | 4 043 | 148,0 | 58 021 |
| 160-179 | 168,0 | 324 | 169,1 | 4 354 | 169,1 | 30 906 |
| 180-199 | 189,0 | 442 | 192,0 | 125 189 | 187,8 | 20 948 |
| 200-299 | 239,0 | 50 462 | 237,6 | 213 733 | 271,9 | 1 026 945 |
| 300-399 | 345,0 | 17 181 | 341,9 | 94 237 | 345,2 | 1 024 164 |
| 400-499 | 447,7 | 11 991 | 441,0 | 61 513 | 445,0 | 616 406 |
| 500-599 | 545,8 | 8 994 | 542,1 | 39 547 | 545,7 | 465 072 |
| 600-699 | 647,8 | 5 048 | 643,1 | 27 100 | 644,2 | 362 290 |
| 700-799 | 746,0 | 3 576 | 742,7 | 18 509 | 744,0 | 278 103 |
| 800-899 | 847,7 | 2 866 | 849,8 | 14 874 | 845,1 | 217 499 |
| 900-999 | 940,9 | 1 597 | 945,8 | 10 287 | 945,3 | 161 653 |
| 1000-1199 | 1 097,6 | 2 790 | 1 082,6 | 18 158 | 1 090,3 | 243 349 |
| 1200-1599 | 1 373,7 | 4 196 | 1 371,1 | 22 601 | 1 373,0 | 269 992 |
| 1600-1999 | 1 758,6 | 2 320 | 1 765,5 | 11 798 | 1 773,4 | 138 991 |
| 2000-2399 | 2 183,8 | 1 633 | 2 178,1 | 6 610 | 2 176,8 | 89 798 |
| 2400 and + | 3 879,3 | 4 393 | 4 149,5 | 23 161 | 4 625,2 | 282 618 |
| $\bar{Y}, \Sigma X_I$ | 603,6 | 119 165 | 557,3 | 706 111 | 811,8 | 5 352 852 |
| CV^2/CV^2+1 | 0,605 | | 0,648 | | 0,598 | |

^aSource: Ministry of Labor - National Department of Labor - 2

^b Industry, commerce and services (excluding unpaid, personal

^c Wages declared to be paid by enterprises to their employees

FORMAL URBAN SECTOR^b
- 1973

| | | | | Brazil | |
|-------------|----------|-------------|---------|-------------|-----------|
| South | | Center West | | | |
| \bar{Y}_I | X_I | \bar{Y}_I | X_I | \bar{Y}_I | X_I |
| 121,8 | 17 993 | 97,4 | 1 494 | 83,2 | 97 068 |
| 149,2 | 9 794 | 151,4 | 1 285 | 148,3 | 73 408 |
| 168,4 | 8 675 | 168,4 | 1 041 | 168,9 | 45 300 |
| 187,9 | 15 109 | 188,6 | 829 | 191,1 | 162 517 |
| 259,0 | 387 901 | 246,1 | 76 100 | 262,8 | 1 755 141 |
| 339,7 | 273 038 | 340,4 | 38 007 | 343,8 | 1 446 627 |
| 444,4 | 136 743 | 442,2 | 28 439 | 444,6 | 855 092 |
| 542,7 | 87 848 | 539,9 | 19 795 | 544,9 | 621 256 |
| 638,7 | 57 471 | 637,0 | 10 658 | 643,3 | 462 567 |
| 743,1 | 39 337 | 741,1 | 7 222 | 743,8 | 346 747 |
| 846,2 | 30 250 | 844,3 | 6 601 | 845,5 | 272 090 |
| 945,7 | 21 411 | 943,6 | 3 808 | 945,3 | 198 756 |
| 1 081,8 | 30 903 | 1 083,7 | 6 778 | 1 088,9 | 301 978 |
| 1 375,2 | 35 701 | 1 374,3 | 6 985 | 1 373,1 | 339 475 |
| 1 770,9 | 17 013 | 1 772,8 | 4 418 | 1 772,4 | 174 540 |
| 2 162,6 | 10 872 | 2 167,4 | 4 872 | 2 175,2 | 113 785 |
| 3 973,6 | 29 033 | 3 635,5 | 11 635 | 4 497,7 | 350 840 |
| 556,6 | 1209 092 | 672,0 | 229 967 | 740,2 | 7 617 187 |
| 0,566 | | 0,582 | | 0,606 | |

/3 Law.
and own - account services).
(Cr\$ 1,00).

relation to the various ranges of salary considered.

The interpretation of these results would suggest immediately the importance of the wages policy adopted. The latter as well as fixing annual increments - to be observed by the public and private sectors - for the workers receiving the minimum salary, also influences directly and proportionally the salary levels of the remaining workers. In this case the results would be simply reflecting the maintenance, more or less constant, of the indices of wages concentration in the period 1970/73.

(b) In terms of the disaggregation of the indices of concentration in the period 1970/73, one observes that the component of wages concentration "between" was not significant - approximately 1.5%. It follows that it was the component of wages concentration "within" regions which really explained the values of the indices encountered.

The southeast region makes the greatest contribution to the concentration of wages observed in the period - about 80% of the values of the indices. This weight is due more to the greater share of its employees in the total for the regions, rather than the value in itself of its index of concentration of wages obtained.

In relative terms, it is worth mentioning that the northeastern, southern and central-west regions were those which showed a growing tendency to concentration of wages in the period. However, wages concentration "within" the southeastern region was highest because of the

greater relative weight of its total of employees.

(c) Finally, it is appropriate to mention the quality of the information used. It is obvious that one cannot take the results obtained as precise indicators of wages concentration occurring in the period 1970/73, because the degree of precision and correctness of the information given by enterprises is not known; furthermore these data do not necessarily refer to the same establishments - which makes the representativeness of the same sample questionable, and finally the fact that the coverage of the information in terms of occupational categories is not known for certain. The most that can be said is that the information refers to those occupational categories which, through individuals, have a professional registration through the issue of work books (C.T.P.S.). These observations certainly restrict any more detailed analysis.

CHAPTER IV

SUMMARY, CONCLUSIONS AND SUGGESTION

The present thesis was divided into three main parts. In the first, three well-known measures of inequality of income were presented and commented on - Pareto's index, Gini's coefficient of concentration and the variance of logarithms. The analysis suggested that among these measures, the last two were the most efficient, both in terms of interpreting the variation of income among individuals (or groups) in the population and in relation to their mathematical/statistical properties.

In the second part, an alternative for understanding the distribution of income was proposed, based on a summary interpretation of the ideas contained in the classic economic theory. From this interpretation, it became possible to elaborate a measure of inequality of income, and also to examine its characteristics, properties and limitations.

Finally, in the third part, the proposed measure was tested in two different cases. On the one hand, it was applied to the distributions of income in Brazil from 1960 to 1970, and on the other to the wage distributions of the employees in the formal urban sector in the period 1970-1973. In this case the implications of the results obtained were commented upon briefly. The most important conclusions were:

- (1) The existing measures of inequality of income have, as a general rule, implicit concepts about the process of income distribution. These concepts can be clearly observed on the basis of the analysis of their operational characteristics and specific

properties. It follows that the proper use of a measure of inequality of income requires the previous verification of its peculiarities in order to ensure compatibility between the hypothesis and the representative empirical results. The ensuing interpretations must necessarily be influenced by the choice of the measures.

(2) The proposed alternative for measuring inequality of income originated in a summary interpretation of the relevant ideas contained in the Classic Economic theory. It was later observed that the Human Capital approach of Neo-Classical theory could also be adopted in applying the measure (see footnote 24). The original interpretation may have been jeopardised in the sense of not being the only one to be put into operational form in the way assumed by the measure. But we demonstrated the possibility of adapting to the measure, other theories dealing with the process of income distribution on the basis of comparable analytical structure.

(3) A principal conclusion of this thesis is that one should try to understand a problem before one starts measuring it. Before choosing or applying a mathematical-statistical model or formula the researcher has to go back to his theory of development, conceptualize the problem, and only then choose or develop a measure which can help him to quantify the relationships which he considers important. In this case, the understanding of the Classics concerning the process of income distribution furnished the development of a measure of inequality with operational characteristics and properties which place it, along with others traditionally known, in

equivalent conditions of application. This is particularly the case of the practical application of the measure suggested in the text, although by this transformation its structural concept of the process of income distribution is simplified rendering it conventional.

Finally, as a suggestion, it is worthwhile highlighting the importance of future studies in this field, especially in the economic circumstances of developing countries, when one observes the insufficiency of traditional models of growth with respect to distribution and rewards generated in the productive process. Its better understanding through the improvement of analytical techniques and methods is fundamental, not only to detect the principal bottlenecks between production and distribution of goods and services, but also for suitable planning of the economic measures attempting to correct the inequality of income arising from the differential ownership of the means of production.

APPENDIX

The proof of the disaggregation of CV^2 in relation to information by income classes (see text) is the following:

$$CV_T^2 = \frac{\sum_{i=1}^X (\bar{Y}_i^2 - \bar{Y}^2)}{X\bar{Y}^2} = \frac{\sum_{i=1}^X (\bar{Y}_i - \bar{Y})^2}{X\bar{Y}^2} = \frac{\sum_{i=1}^X \left[(\bar{Y}_i - \bar{Y}_I) + (\bar{Y}_I - \bar{Y}) \right]^2}{X\bar{Y}^2} =$$

$$= \frac{\sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I)^2 + \sum_{i=1}^X (\bar{Y}_I - \bar{Y})^2 - 2 \sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I) (\bar{Y}_I - \bar{Y})}{X\bar{Y}^2}$$

For any class of income I , the term $(\bar{Y}_I - \bar{Y})$ is the same for all the values of i , hence the last term of the above expression reduces to:

$$2 \sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I) \cdot (\bar{Y}_I - \bar{Y}) = 2 \sum_{I=1}^K (\bar{Y}_I - \bar{Y}) \cdot \sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I)$$

and for any class of income I ,

$$\sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I) = \sum_{i=1}^X \bar{Y}_i - \sum_{i=1}^X \bar{Y}_I = \sum_{i=1}^X \bar{Y}_i - \sum_{I=1}^K X_I \bar{Y}_I = 0$$

consequently,

$$2 \sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I) \cdot (\bar{Y}_I - \bar{Y}) = 0$$

Finally,

$$CV_T^2 = \frac{\sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I)^2 + (\bar{Y}_I - \bar{Y})^2}{XY}$$

but, as

$$\sum_{i=1}^X i = \sum_{I=1}^K X_I = X$$

we have

$$CV_T^2 = \frac{\sum_{I=1}^K X_I \bar{Y}_I^2}{XY^2} \cdot \frac{\sum_{i=1}^X (\bar{Y}_i - \bar{Y}_I)^2}{X_I \bar{Y}_I^2} + \frac{\sum_{I=1}^K (\bar{Y}_I - \bar{Y})^2}{X \bar{Y}^2},$$

or simply,

$$CV_T^2 = \frac{\sum_{I=1}^K X_I \bar{Y}_I^2}{XY^2} \cdot (CV_W^2) + CV_B^2.$$

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